# Lecture 1 - Part 2 Some Basics 

(January 9, 2015)

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## Regression

- training data $\left\{\left(\boldsymbol{x}_{i}, y_{i}\right) ; i=1,2, \ldots, n\right\}$
- $\boldsymbol{x}_{i} \in \mathbb{R}^{d}$ and $y_{i} \in \mathbb{R}$
- want $f(\cdot)$ so that we can predict $y$ from $\boldsymbol{x}$ with $f(\boldsymbol{x})$
- mean squared error,

$$
\operatorname{MSE}(f)=\mathbb{E}[y-f(\boldsymbol{x})]^{2},
$$

a common criterion for how good $f$ is

## $\mathbb{E}(y \mid \boldsymbol{x})$

Let $g(\boldsymbol{x})=\mathbb{E}(y \mid \boldsymbol{x})$, and $h(\boldsymbol{x})$ be any other function of $\boldsymbol{x}$. Then,

$$
\begin{aligned}
& \mathbb{E}[y-h(\boldsymbol{x})]^{2} \\
= & \mathbb{E}[y-g(\boldsymbol{x})+g(\boldsymbol{x})-h(\boldsymbol{x})]^{2} \\
= & \mathbb{E}[y-g(\boldsymbol{x})]^{2}+\underbrace{\mathbb{E}[g(\boldsymbol{x})-h(\boldsymbol{x})]^{2}}_{\geq 0}+2 \underbrace{\mathbb{E}[(y-g(\boldsymbol{x}))(g(\boldsymbol{x})-h(\boldsymbol{x}))]}_{=0}
\end{aligned}
$$

$$
\geq \mathbb{E}[y-g(\boldsymbol{x})]^{2}
$$

## $\underline{\text { main task for regression: "go after" the function, } \mathbb{E}(y \mid x)}$

Exercise Show that $\mathbb{E}[(y-g(\boldsymbol{x}))(g(\boldsymbol{x})-h(\boldsymbol{x}))]=0$. (Hint: Use $\mathbb{E}(\cdot)=\mathbb{E}[\mathbb{E}(\cdot \mid \boldsymbol{x})]$.

## Ex I: Linear Regression

- for $x \in \mathbb{R}$, can start by modelling $\mathbb{E}(y \mid x)$ as

$$
f(x)=\alpha+\beta x
$$

- further justification: $\mathbb{E}(y \mid x)$ linear in $x$ if $(x, y)$ have a joint normal distribution
- just have to estimate $\alpha$ and $\beta$ from training data
- for $\boldsymbol{x} \in \mathbb{R}^{d}$, simply

$$
f(\boldsymbol{x})=\alpha+\boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}
$$

## Ex II: Nearest-Neighbor Regression

- may feel uncomfortable with assuming $\mathbb{E}(y \mid \boldsymbol{x})$ to be linear
- can choose to estimate $\mathbb{E}(y \mid \boldsymbol{x})$ by

$$
\widehat{\mathbb{E}}(y \mid \boldsymbol{x})=\text { average }\left\{y_{i}: \boldsymbol{x}_{i} \in \mathcal{N}(\boldsymbol{x})\right\}
$$

where $\mathcal{N}(\boldsymbol{x})$ denotes a "neighborhood" around $\boldsymbol{x}$

- meaning of $\mathbb{E}(y \mid \boldsymbol{x})$ : the average of $y$ given a particular $\boldsymbol{x}$
- almost literal interpretation of $\mathbb{E}(y \mid \boldsymbol{x})$
- relaxes "given a particular $x$ " to "within a neighborhood of $x$ "


## Ex II: Nearest-Neighbor Regression

- for $x \in \mathbb{R}$, suppose

$$
\mathcal{N}(x)=\left\{x_{i}: \frac{\left|x_{i}-x\right|}{h}<1\right\}
$$

- can express estimate as

$$
\widehat{\mathbb{E}}(y \mid x)=\frac{\sum_{i=1}^{n} I\left(\frac{\left|x_{i}-x\right|}{h}<1\right) y_{i}}{\sum_{i=1}^{n} I\left(\frac{\left|x_{i}-x\right|}{h}<1\right)}
$$

- need to specify $h$ a priori ... called a tuning parameter


## Ex III: Kernel Regression

- can further generalize to

$$
\widehat{\mathbb{E}}(y \mid x)=\frac{\sum_{i=1}^{n} \frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right) y_{i}}{\sum_{i=1}^{n} \frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right)}
$$

where $K(u)$ is a kernel function such that

$$
\int K(u) d u=1, \quad \int u K(u) d u=0, \quad \int u^{2} K(u) d u<\infty
$$

- simple average vs weighted average

(a) $K(u)=\frac{1}{2} I(|u|<1) ;(\mathrm{b}) K(u)=\frac{1}{\sqrt{2 \pi}} e^{-u^{2} / 2}$.


## Bias-Variance Analysis

- typical model assumption:

$$
\begin{gathered}
y_{i}=f\left(x_{i}\right)+\varepsilon_{i} \\
\mathbb{E}\left(\varepsilon_{i}\right)=0, \quad \operatorname{Var}\left(\varepsilon_{i}\right)=\sigma^{2}, \quad \operatorname{Cov}\left(\varepsilon_{i}, \varepsilon_{j}\right)=0 \text { for } i \neq j
\end{gathered}
$$

- let

$$
w_{i}=\frac{\frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right)}{\sum_{j=1}^{n} \frac{1}{n h} K\left(\frac{x_{j}-x}{h}\right)} \quad \text { so that } \quad \widehat{f}(x)=\sum_{i=1}^{n} w_{i} y_{i}
$$

- further simplifying assumption: $x_{i} \sim \operatorname{Unif}(0,1)$


## Bias

$$
\begin{gathered}
\widehat{f}(x)=\sum w_{i} y_{i} \Rightarrow \mathbb{E}[\widehat{f}(x)]=\sum w_{i} \mathbb{E}\left(y_{i}\right)=\sum w_{i} f\left(x_{i}\right) \\
f\left(x_{i}\right) \approx f(x)+\left(x_{i}-x\right) f^{\prime}(x)+\frac{1}{2}\left(x_{i}-x\right)^{2} f^{\prime \prime}(x)
\end{gathered}
$$

$$
\mathbb{E}[\widehat{f}(x)] \approx
$$

$$
f(x) \underbrace{\sum w_{i}}_{=1}+f^{\prime}(x) \underbrace{\sum w_{i}\left(x_{i}-x\right)}_{\approx 0}+\frac{1}{2} f^{\prime \prime}(x) \underbrace{\sum w_{i}\left(x_{i}-x\right)^{2}}_{\approx h^{2}\left[\int u^{2} K(u) d u\right]}
$$

$\therefore \quad \operatorname{Bias}[\widehat{f}(x)] \approx h^{2} B_{0} \quad$ where $\quad B_{0}=\frac{1}{2} f^{\prime \prime}(x)\left[\int u^{2} K(u) d u\right]$

## Variance

$$
\begin{aligned}
& \widehat{f}(x)=\sum w_{i} y_{i} \Rightarrow \\
& \mathbb{V} \operatorname{ar}[\widehat{f}(x)]=\sum w_{i}^{2} \mathbb{V} \operatorname{ar}\left(y_{i}\right)=\sigma^{2} \underbrace{\left[\sum w_{i}^{2}\right]}_{\approx \frac{1}{n h}\left[\int K^{2}(u) d u\right]} \\
& \therefore \quad \operatorname{Var}[\widehat{f}(x)] \approx \frac{1}{n h} V_{0} \quad \text { where } \quad V_{0}=\sigma^{2}\left[\int K^{2}(u) d u\right]
\end{aligned}
$$

## Discussion

$$
\begin{aligned}
& h \uparrow \Rightarrow \text { bias } \uparrow \text { and variance } \downarrow \\
& h \downarrow \Rightarrow \text { bias } \downarrow \text { and variance } \uparrow
\end{aligned}
$$

Are these intuitively "obvious"?

## Some Details

First,

$$
\sum w_{i}\left(x_{i}-x\right)=\frac{\sum \frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right)\left(x_{i}-x\right)}{\sum \frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right)}
$$

where

$$
\text { numerator } \approx \int\left(\frac{v-x}{h}\right) K\left(\frac{v-x}{h}\right) d v \stackrel{(*)}{=} h \int u K(u) d u=0
$$

and

$$
\text { denominator } \approx \int \frac{1}{h} K\left(\frac{v-x}{h}\right) d v \stackrel{(*)}{=} \int K(u) d u=1
$$

$\left.{ }^{*}\right) u=(v-x) / h, d u=(1 / h) d v$

## Some Details

Likewise,

$$
\sum w_{i}\left(x_{i}-x\right)^{2}=\frac{\sum \frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right)\left(x_{i}-x\right)^{2}}{\sum \frac{1}{n h} K\left(\frac{x_{i}-x}{h}\right)}
$$

where

$$
\text { denominator } \approx 1 \quad \text { (previous slide })
$$

and

$$
\text { numerator } \approx \int h\left(\frac{v-x}{h}\right)^{2} K\left(\frac{v-x}{h}\right) d v=h^{2} \int u^{2} K(u) d u
$$

## Some Details

Exercise Use the same argument to "show" that

$$
\sum w_{i}^{2} \approx \frac{1}{n h}\left[\int K^{2}(u) d u\right]
$$

## Bias-Variance Trade-Off

$$
\begin{aligned}
\operatorname{MSE}(\widehat{f}) & \equiv \mathbb{E}(\widehat{f}-f)^{2} \\
& =\mathbb{E}[\widehat{f}-\mathbb{E}(\widehat{f})+\mathbb{E}(\widehat{f})-f]^{2} \\
& =\underbrace{\mathbb{E}[\widehat{f}-\mathbb{E}(\widehat{f})]^{2}}_{\operatorname{Var}(\widehat{f})}+\underbrace{[\mathbb{E}(\widehat{f})-f]^{2}}_{\operatorname{Bias}^{2}(\widehat{f})}+2 \underbrace{\mathbb{E}[(\widehat{f}-\mathbb{E}(\widehat{f}))(\mathbb{E}(\widehat{f})-f)]}_{=0}
\end{aligned}
$$

Exercise Show that $\mathbb{E}[(\widehat{f}-\mathbb{E}(\widehat{f}))(\mathbb{E}(\widehat{f})-f)]=0$.

## Bias-Variance Trade-Off

- for kernel regression,

$$
\mathrm{MSE}=\operatorname{Var}+\operatorname{Bias}^{2} \approx h^{4} B_{0}^{2}+\frac{V_{0}}{n h}
$$

- can find the "optimal" $h$ (in terms of the MSE):

$$
\frac{d}{d h} \mathrm{MSE} \approx 4 B_{0}^{2} h^{3}-\frac{V_{0}}{n h^{2}}=0 \quad \Rightarrow \quad h^{*} \sim O\left(n^{-1 / 5}\right)
$$

- general phenomenon, not just for kernel regression


## Curse of Dimensionality

For $\boldsymbol{x} \in \mathbb{R}^{d}$, neighborhood-based methods such as kernel regression still apply ("just" use $K(\boldsymbol{u})$ for $\boldsymbol{u} \in \mathbb{R}^{d}$ ) but they become increasingly difficult.

Example Suppose data are uniformly distributed inside the unit ball, $\{\boldsymbol{x}:\|\boldsymbol{x}\| \leq 1\}$. Consider a neighborhood around $\mathbf{0}$ with radius $h<1$. What fraction of data does the neighborhood contain?

$$
(\text { fraction of data })=\frac{\operatorname{vol}(\text { neighborhood })}{\operatorname{vol}(\text { unit ball })}=\frac{\frac{\pi^{d / 2}}{\Gamma(d / 2+1)} h^{d}}{\frac{\pi^{d / 2}}{\Gamma(d / 2+1)} 1^{d}}=h^{d} .
$$

Thus, in $d=100$ dimensions, even a neighborhood with radius $h=0.95$ contains $<0.6 \%$ of the data.

## Classification

- training data $\left\{\left(\boldsymbol{x}_{i}, y_{i}\right) ; i=1,2, \ldots, n\right\}$
- $\boldsymbol{x}_{i} \in \mathbb{R}^{d}$ and $y_{i} \in\{1,2, \ldots, K\}$
- want $f(\cdot)$ so that we can classify $y$ from $\boldsymbol{x}$ with $f(\boldsymbol{x})$
- mean 0-1 error,

$$
\operatorname{error}(f)=\mathbb{E}[I(y \neq f(\boldsymbol{x}))],
$$

a common criterion for how good $f$ is

$$
\mathbb{P}(y \mid \boldsymbol{x})
$$

Exercise Show that the function that minimizes the mean 0-1 error is

$$
f(\boldsymbol{x})=\underset{k=1, \ldots, K}{\arg \max } \mathbb{P}(y=k \mid \boldsymbol{x}) .
$$

$\Downarrow$ $\underline{\text { main task for classification: "go after" the function, } \mathbb{P}(y \mid x)}$

Remark For binary $y \in\{0,1\}$, also have $\mathbb{E}(y \mid \boldsymbol{x})=\mathbb{P}(y=1 \mid \boldsymbol{x})$.

## Two Strategies

- "go after" $\mathbb{P}(y \mid \boldsymbol{x}) \underline{\text { directly }}$
- use Bayes theorem,

$$
\mathbb{P}(y=k \mid \boldsymbol{x})=\frac{\pi_{k} p_{k}(\boldsymbol{x})}{\pi_{1} p_{1}(\boldsymbol{x})+\ldots+\pi_{K} p_{K}(\boldsymbol{x})}
$$

and "go after" $\mathbb{P}(y \mid \boldsymbol{x})$ indirectly by first "going after"

- $p_{k}(\boldsymbol{x})$, the conditional distribution of $\boldsymbol{x} \mid \boldsymbol{y}=k$, and
$-\pi_{k}$, the prior probability of class $k$,
for each $k=1, \ldots, K$


## Ex IV: Logistic Regression

- for binary $y \in\{0,1\}$, can model

$$
\mathbb{P}(y=1 \mid \boldsymbol{x})=\frac{\exp \left\{\alpha+\boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}\right\}}{1+\exp \left\{\alpha+\boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}\right\}}
$$

or, equivalently,

$$
\log \frac{\mathbb{P}(y=1 \mid \boldsymbol{x})}{\mathbb{P}(y=0 \mid \boldsymbol{x})}=\alpha+\boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}
$$

- just have to estimate $\alpha$ and $\boldsymbol{\beta}$ from training data


## Ex V: Linear Discriminant Analysis

- alternatively, can model

$$
p_{k}(\boldsymbol{x}) \sim \mathrm{N}\left(\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}\right), \quad k=0,1
$$

- recall multivariate normal density function

$$
p_{k}(\boldsymbol{x})=\frac{1}{\sqrt{(2 \pi)^{d}|\boldsymbol{\Sigma}|}} \exp \left[-\frac{1}{2}\left(\boldsymbol{x}-\boldsymbol{\mu}_{k}\right)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}\left(\boldsymbol{x}-\boldsymbol{\mu}_{k}\right)\right]
$$

## Ex V: Linear Discriminant Analysis

- then,

$$
\log \frac{\mathbb{P}(y=1 \mid \boldsymbol{x})}{\mathbb{P}(y=0 \mid \boldsymbol{x})}==\log \frac{\pi_{1}}{\pi_{0}}+\log \frac{p_{1}(\boldsymbol{x})}{p_{0}(\boldsymbol{x})}
$$

where

$$
\begin{aligned}
\log \frac{p_{1}(\boldsymbol{x})}{p_{0}(\boldsymbol{x})}=-\frac{1}{2}\left[\left(\boldsymbol{x}-\boldsymbol{\mu}_{1}\right)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}( \right. & \left.\boldsymbol{x}-\boldsymbol{\mu}_{1}\right) \\
& \left.-\left(\boldsymbol{x}-\boldsymbol{\mu}_{0}\right)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}\left(\boldsymbol{x}-\boldsymbol{\mu}_{0}\right)\right]
\end{aligned}
$$

- just have to estimate $\pi_{1}, \pi_{0}, \boldsymbol{\mu}_{1}, \boldsymbol{\mu}_{0}$ and $\boldsymbol{\Sigma}\left(\right.$ actually, $\left.\boldsymbol{\Sigma}^{-1}\right)$


## Comparison

- slight rearrangement of the last equation from the previous slide gives

$$
\begin{aligned}
& \log \frac{\mathbb{P}(y=1 \mid \boldsymbol{x})}{\mathbb{P}(y=0 \mid \boldsymbol{x})}=\underbrace{\log \frac{\pi_{1}}{\pi_{0}}-\frac{1}{2}\left(\boldsymbol{\mu}_{1}^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_{1}-\boldsymbol{\mu}_{0}^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_{0}\right)}_{\alpha} \\
&+\underbrace{\left(\boldsymbol{\mu}_{1}-\boldsymbol{\mu}_{0}\right)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} \boldsymbol{x}}_{\boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}}
\end{aligned}
$$

- linear discriminant analysis and logistic regression: two different ways of "going after" the same, linear decision boundary


## Which Is Better?

B. Efron (1975), "The efficiency of logistic regression compared to normal discriminant analysis", JASA 70, pp. 892-898.

- $n \rightarrow \infty$; fixed $d$
- if $p_{k}$ normal, logistic regression less efficient
- loss of efficiency between $1 / 3$ to $1 / 2$


## Enter "Big Data"

- if all this sounds easy, don't forget $\boldsymbol{\Sigma}$ is $d \times d$
- very hard for relatively large $d$
- $\boldsymbol{\Sigma}^{-1}$ can be estimated by the graphical LASSO (Friedman, Hastie \& Tibshirani, 2008; Biostatistics)
- Fan, Feng \& Wu (2009; Ann. Appl. Stat.) applied gLASSO-estimated $\boldsymbol{\Sigma}^{-1}$ to perform linear discriminant analysis
- Cai \& Liu (2012; JASA) proposed to estimate $\boldsymbol{\beta} \equiv \boldsymbol{\Sigma}^{-1}\left(\boldsymbol{\mu}_{1}-\boldsymbol{\mu}_{0}\right)$ directly with sparsity constraints

Research Perform an analysis like that of Efron (1975) when $d \rightarrow \infty$ as well.

## Ex VI: Naïve Bayes

- may feel uncomfortable with assuming $p_{k}(\boldsymbol{x}) \sim \mathrm{N}\left(\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}\right)$
- data scientists have long used (and still like) the model,

$$
p_{k}(\boldsymbol{x})=\prod_{j=1}^{d} f_{k, j}\left(x_{j}\right)
$$

where each $f_{k, j}(\cdot)$ can be estimated separately

- especially helpful if the predictors are of mixed types (e.g., some continuous, some categorical)


## Ex VI: Naïve Bayes

- may feel uncomfortable with assuming independence
- but

$$
\begin{aligned}
\log \frac{\mathbb{P}(y=1 \mid \boldsymbol{x})}{\mathbb{P}(y=0 \mid \boldsymbol{x})} & =\underbrace{\log \frac{\pi_{1}}{\pi_{0}}}_{\alpha}+\sum_{j=1}^{d} \underbrace{\log \frac{f_{1, j}\left(x_{j}\right)}{f_{0, j}\left(x_{j}\right)}}_{g_{j}\left(x_{j}\right)} \\
& \equiv \alpha+\sum_{j=1}^{d} g_{j}\left(x_{j}\right),
\end{aligned}
$$

and most people comfortable with generalizing linear logistic regression to additive logistic regression

## Ex VII: Neural Networks

- sigmoid function

$$
\sigma(u)=\frac{e^{u}}{1+e^{u}}
$$

- hidden layer $\ell=1,2, \ldots, L-1$

$$
z_{t}^{(\ell)}=\sigma\left(\alpha_{t}^{(\ell)}+\sum_{b} w_{t, b}^{(\ell)} z_{b}^{(\ell-1)}\right)
$$

- top layer $L$

$$
z_{t}^{(L)}=\mathbb{P}(y=t \mid \ldots)
$$

- bottom layer 0

$$
z_{b}^{(0)}=x_{b}, \quad b=1,2, \ldots, d
$$

## Ex VIII: Nearest-Neighbor Classifier

- can also estimate $\mathbb{P}(y \mid \boldsymbol{x})$ by

$$
\widehat{\mathbb{P}}(y=k \mid \boldsymbol{x})=\text { fraction }\left\{y_{i}=k: \boldsymbol{x}_{i} \in \mathcal{N}(\boldsymbol{x})\right\},
$$

where $\mathcal{N}(\boldsymbol{x})$ denotes a "neighborhood" around $\boldsymbol{x}$

## Bayes Error

Myth If my misclassification error is not very close to zero, my classifier must not be very good.

Myth If I know the true model, my misclassification error must be zero.

Truth Even if we knew $\mathbb{P}(y \mid \boldsymbol{x})$ (or $p_{k}(\boldsymbol{x}), \pi_{k}, \ldots$ ) perfectly, we might still have considerable misclassification error - these errors are called the Bayes error.

## Bayes Error

$$
\begin{gathered}
x \in \mathbb{R} \\
\pi_{1}=\pi_{0}=1 / 2 \\
p_{1}(x) \sim \mathrm{N}\left(\mu_{1}, \sigma^{2}\right) \\
p_{0}(x) \sim \mathrm{N}\left(\mu_{0}, \sigma^{2}\right)
\end{gathered}
$$



Exercise How does the Bayes error change with $\Delta \equiv\left|\mu_{1}-\mu_{0}\right|$, and with $\sigma^{2}$ ?

Question Can we reduce the Bayes error?

## Summary

- key ideas:
- regression; mean squared error; $\mathbb{E}(y \mid \boldsymbol{x})$
- bias-variance trade-off; curse of dimensionality
- classification; mean 0-1 error; $\mathbb{P}(y \mid \boldsymbol{x})$; Bayes error
- specific methods:
- linear regression; nearest-neighbors; kernel regression
- logistic regression; linear discriminant analysis
- naïve Bayes; neural network
- graphical LASSO
- didn't discuss:
- actual estimation procedures


## Next ...

- course administration, logistics, etc

