Lecture 1 - Part 2**Some Basics** (January 9, 2015) Mu Zhu University of Waterloo

Regression

- training data $\{(x_i, y_i); i = 1, 2, ..., n\}$
- $\boldsymbol{x}_i \in \mathbb{R}^d$ and $\boldsymbol{y}_i \in \mathbb{R}$
- want $f(\cdot)$ so that we can predict y from x with f(x)
- mean squared error,

$$MSE(f) = \mathbb{E}[y - f(\boldsymbol{x})]^2,$$

a common criterion for how good f is

$$\mathbb{E}(y|\boldsymbol{x})$$

Let $g(\boldsymbol{x}) = \mathbb{E}(y|\boldsymbol{x})$, and $h(\boldsymbol{x})$ be any other function of \boldsymbol{x} . Then, $\mathbb{E}[y-h(\boldsymbol{x})]^2$ $= \mathbb{E}[y - q(\boldsymbol{x}) + q(\boldsymbol{x}) - h(\boldsymbol{x})]^2$ $= \mathbb{E}[y - g(\boldsymbol{x})]^2 + \underbrace{\mathbb{E}[g(\boldsymbol{x}) - h(\boldsymbol{x})]^2}_{=} + 2\underbrace{\mathbb{E}[(y - g(\boldsymbol{x}))(g(\boldsymbol{x}) - h(\boldsymbol{x}))]}_{=}$ >0=0 $\geq \mathbb{E}[y-q(\boldsymbol{x})]^2.$ \Downarrow main task for regression: "go after" the function, $\mathbb{E}(y|\boldsymbol{x})$ **Exercise** Show that $\mathbb{E}[(y - g(\boldsymbol{x}))(g(\boldsymbol{x}) - h(\boldsymbol{x}))] = 0$. (<u>Hint</u>: Use $\mathbb{E}(\cdot) = \mathbb{E}[\mathbb{E}(\cdot|\boldsymbol{x})].)$

Ex I: Linear Regression

• for $x \in \mathbb{R}$, can start by modelling $\mathbb{E}(y|x)$ as

$$f(x) = \alpha + \beta x$$

- further justification: $\mathbb{E}(y|x)$ linear in x if (x, y) have a joint normal distribution
- just have to estimate α and β from training data
- for $\boldsymbol{x} \in \mathbb{R}^d$, simply

$$f(\boldsymbol{x}) = \alpha + \boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}$$

Ex II: Nearest-Neighbor Regression

- may feel uncomfortable with assuming $\mathbb{E}(y|\boldsymbol{x})$ to be linear
- can choose to estimate $\mathbb{E}(y|\boldsymbol{x})$ by

 $\widehat{\mathbb{E}}(y|\boldsymbol{x}) = \operatorname{average} \{y_i : \boldsymbol{x}_i \in \mathcal{N}(\boldsymbol{x})\},\$

where $\mathcal{N}(\boldsymbol{x})$ denotes a "neighborhood" around \boldsymbol{x}

- meaning of $\mathbb{E}(y|\boldsymbol{x})$: the average of y given a particular \boldsymbol{x}
- almost literal interpretation of $\mathbb{E}(y|\boldsymbol{x})$
- relaxes "given a particular x" to "within a neighborhood of x"

Ex II: Nearest-Neighbor Regression

• for $x \in \mathbb{R}$, suppose

$$\mathcal{N}(x) = \left\{ x_i : \frac{|x_i - x|}{h} < 1 \right\}$$

• can express estimate as

$$\widehat{\mathbb{E}}(y|x) = \frac{\sum_{i=1}^{n} I\left(\frac{|x_i - x|}{h} < 1\right) y_i}{\sum_{i=1}^{n} I\left(\frac{|x_i - x|}{h} < 1\right)}$$

• need to specify h a priori ... called a tuning parameter

Ex III: Kernel Regression

• can further generalize to

$$\widehat{\mathbb{E}}(y|x) = \frac{\sum_{i=1}^{n} \frac{1}{nh} K\left(\frac{x_i - x}{h}\right) y_i}{\sum_{i=1}^{n} \frac{1}{nh} K\left(\frac{x_i - x}{h}\right)},$$

where K(u) is a kernel function such that

$$\int K(u)du = 1, \quad \int uK(u)du = 0, \quad \int u^2 K(u)du < \infty$$

• simple average vs weighted average



Bias-Variance Analysis

• typical model assumption:

$$y_i = f(x_i) + \varepsilon_i$$

 $\mathbb{E}(\varepsilon_i) = 0, \quad \mathbb{V}ar(\varepsilon_i) = \sigma^2, \quad \operatorname{Cov}(\varepsilon_i, \varepsilon_j) = 0 \text{ for } i \neq j$

• let

$$w_{i} = \frac{\frac{1}{nh}K\left(\frac{x_{i} - x}{h}\right)}{\sum_{j=1}^{n}\frac{1}{nh}K\left(\frac{x_{j} - x}{h}\right)} \quad \text{so that} \quad \widehat{f}(x) = \sum_{i=1}^{n}w_{i}y_{i}$$

• further simplifying assumption: $x_i \sim \text{Unif}(0, 1)$

$$\mathbf{Bias}$$

$$\widehat{f}(x) = \sum w_i y_i \quad \Rightarrow \quad \mathbb{E}\left[\widehat{f}(x)\right] = \sum w_i \mathbb{E}(y_i) = \sum w_i f(x_i)$$

$$f(x_i) \approx f(x) + (x_i - x)f'(x) + \frac{1}{2}(x_i - x)^2 f''(x)$$

$$\Downarrow$$

$$\mathbb{E}\left[\widehat{f}(x)\right] \approx$$

$$f(x) \sum_{i=1}^{\infty} w_i + f'(x) \sum_{i=0}^{\infty} w_i (x_i - x) + \frac{1}{2} f''(x) \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i (x_i - x) + \frac{1}{2} f''(x) \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i (x_i - x) + \frac{1}{2} f''(x) \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i (x_i - x) + \frac{1}{2} f''(x) \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i (x_i - x) + \frac{1}{2} f''(x) \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i (x_i - x) + \frac{1}{2} f''(x) \sum_{i=1}^{\infty} w_i (x_i - x)^2 \sum_{i=1}^{\infty} w_i ($$

$$\widehat{f}(x) = \sum w_i y_i \Rightarrow$$
 $\mathbb{V}ar\left[\widehat{f}(x)\right] = \sum w_i^2 \mathbb{V}ar(y_i) = \sigma^2 \underbrace{\left[\sum w_i^2\right]}_{\approx \frac{1}{nh} [\int K^2(u) du]}$
 $\therefore \mathbb{V}ar\left[\widehat{f}(x)\right] \approx \frac{1}{nh} V_0$ where $V_0 = \sigma^2 \left[\int K^2(u) du\right]$



Some Details

First,

$$\sum w_i(x_i - x) = \frac{\sum \frac{1}{nh} K\left(\frac{x_i - x}{h}\right) (x_i - x)}{\sum \frac{1}{nh} K\left(\frac{x_i - x}{h}\right)},$$

where

numerator
$$\approx \int \left(\frac{v-x}{h}\right) K\left(\frac{v-x}{h}\right) dv \stackrel{(*)}{=} h \int u K(u) du = 0,$$

and

denominator
$$\approx \int \frac{1}{h} K\left(\frac{v-x}{h}\right) dv \stackrel{(*)}{=} \int K(u) du = 1.$$

(*)
$$u = (v - x)/h, du = (1/h)dv$$

Some Details

Likewise,

$$\sum w_i (x_i - x)^2 = \frac{\sum \frac{1}{nh} K\left(\frac{x_i - x}{h}\right) (x_i - x)^2}{\sum \frac{1}{nh} K\left(\frac{x_i - x}{h}\right)},$$

where

denominator ≈ 1 (previous slide)

and

numerator
$$\approx \int h\left(\frac{v-x}{h}\right)^2 K\left(\frac{v-x}{h}\right) dv = h^2 \int u^2 K(u) du.$$



Bias-Variance Trade-Off

$$\begin{split} \mathrm{MSE}(\widehat{f}) &\equiv & \mathbb{E}(\widehat{f} - f)^2 \\ &= & \mathbb{E}[\widehat{f} - \mathbb{E}(\widehat{f}) + \mathbb{E}(\widehat{f}) - f]^2 \\ &= & \underbrace{\mathbb{E}[\widehat{f} - \mathbb{E}(\widehat{f})]^2}_{\mathrm{Var}(\widehat{f})} + \underbrace{[\mathbb{E}(\widehat{f}) - f]^2}_{\mathrm{Bias}^2(\widehat{f})} + 2\underbrace{\mathbb{E}[(\widehat{f} - \mathbb{E}(\widehat{f}))(\mathbb{E}(\widehat{f}) - f)]}_{=0} \end{split}$$

Exercise Show that $\mathbb{E}[(\widehat{f} - \mathbb{E}(\widehat{f}))(\mathbb{E}(\widehat{f}) - f)] = 0.$

Bias-Variance Trade-Off

• for kernel regression,

$$MSE = Var + Bias^2 \approx h^4 B_0^2 + \frac{V_0}{nh}$$

• can find the "optimal" h (in terms of the MSE):

$$\frac{d}{dh} \text{MSE} \approx 4B_0^2 h^3 - \frac{V_0}{nh^2} = 0 \quad \Rightarrow \quad h^* \sim O(n^{-1/5})$$

• general phenomenon, not just for kernel regression

Curse of Dimensionality

For $\boldsymbol{x} \in \mathbb{R}^d$, neighborhood-based methods such as kernel regression still apply ("just" use $K(\boldsymbol{u})$ for $\boldsymbol{u} \in \mathbb{R}^d$) but they become increasingly difficult.

Example Suppose data are uniformly distributed inside the unit ball, $\{x : ||x|| \le 1\}$. Consider a neighborhood around **0** with radius h < 1. What fraction of data does the neighborhood contain?

(fraction of data) =
$$\frac{\text{vol(neighborhood)}}{\text{vol(unit ball)}} = \frac{\frac{\pi^{d/2}}{\Gamma(d/2+1)}h^d}{\frac{\pi^{d/2}}{\Gamma(d/2+1)}1^d} = h^d.$$

Thus, in d = 100 dimensions, even a neighborhood with radius h = 0.95 contains < 0.6% of the data.

Classification

• training data $\{(x_i, y_i); i = 1, 2, ..., n\}$

•
$$\boldsymbol{x}_i \in \mathbb{R}^d \text{ and } \boldsymbol{y}_i \in \{1, 2, ..., K\}$$

- want $f(\cdot)$ so that we can classify y from x with f(x)
- mean 0-1 error,

$$\operatorname{error}(f) = \mathbb{E}[I(y \neq f(\boldsymbol{x}))],$$

a common criterion for how good f is



Two Strategies

- "go after" $\mathbb{P}(y|\boldsymbol{x})$ directly
- use Bayes theorem,

$$\mathbb{P}(y = k | \boldsymbol{x}) = \frac{\pi_k p_k(\boldsymbol{x})}{\pi_1 p_1(\boldsymbol{x}) + \ldots + \pi_K p_K(\boldsymbol{x})},$$

and "go after" $\mathbb{P}(y|\boldsymbol{x})$ indirectly by first "going after"

 $-p_k(\boldsymbol{x})$, the conditional distribution of $\boldsymbol{x}|y=k$, and

 $-\pi_k$, the prior probability of class k,

for each k = 1, ..., K

Ex IV: Logistic Regression

• for binary $y \in \{0, 1\}$, can model

$$\mathbb{P}(y=1|\boldsymbol{x}) = \frac{\exp\{\alpha + \boldsymbol{\beta}^{\mathrm{T}}\boldsymbol{x}\}}{1 + \exp\{\alpha + \boldsymbol{\beta}^{\mathrm{T}}\boldsymbol{x}\}},$$

or, equivalently,

$$\log \frac{\mathbb{P}(y=1|\boldsymbol{x})}{\mathbb{P}(y=0|\boldsymbol{x})} = \alpha + \boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}$$

• just have to estimate α and β from training data

Ex V: Linear Discriminant Analysis

• alternatively, can model

$$p_k(\boldsymbol{x}) \sim \mathrm{N}(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}), \quad k = 0, 1$$

• recall multivariate normal density function

$$p_k(\boldsymbol{x}) = \frac{1}{\sqrt{(2\pi)^d |\boldsymbol{\Sigma}|}} \exp\left[-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}(\boldsymbol{x} - \boldsymbol{\mu}_k)\right]$$

Ex V: Linear Discriminant Analysis

• then,

$$\log \frac{\mathbb{P}(y=1|\boldsymbol{x})}{\mathbb{P}(y=0|\boldsymbol{x})} = \log \frac{\pi_1}{\pi_0} + \log \frac{p_1(\boldsymbol{x})}{p_0(\boldsymbol{x})}$$

where

$$\log \frac{p_1(\boldsymbol{x})}{p_0(\boldsymbol{x})} = -\frac{1}{2} [(\boldsymbol{x} - \boldsymbol{\mu}_1)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu}_1) - (\boldsymbol{x} - \boldsymbol{\mu}_0)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu}_0)]$$

• just have to estimate $\pi_1, \pi_0, \mu_1, \mu_0$ and Σ (actually, Σ^{-1})

Comparison

• slight rearrangement of the last equation from the previous slide gives

$$\log \frac{\mathbb{P}(y=1|\boldsymbol{x})}{\mathbb{P}(y=0|\boldsymbol{x})} = \underbrace{\log \frac{\pi_1}{\pi_0} - \frac{1}{2} \left(\boldsymbol{\mu}_1^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_1 - \boldsymbol{\mu}_0^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_0\right)}_{\alpha} + \underbrace{(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0)^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} \boldsymbol{x}}_{\boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}}$$

• linear discriminant analysis and logistic regression: two different ways of "going after" the same, linear decision boundary

Which Is Better?

B. Efron (1975), "The efficiency of logistic regression compared to normal discriminant analysis", *JASA* **70**, pp. 892–898.

- $n \to \infty$; fixed d
- if p_k normal, logistic regression less efficient
- loss of efficiency between 1/3 to 1/2

Enter "Big Data"

- if all this sounds easy, don't forget Σ is $d \times d$
- very hard for relatively large d
- Σ^{-1} can be estimated by the graphical LASSO (Friedman, Hastie & Tibshirani, 2008; *Biostatistics*)
- Fan, Feng & Wu (2009; Ann. Appl. Stat.) applied gLASSO-estimated Σ^{-1} to perform linear discriminant analysis
- Cai & Liu (2012; *JASA*) proposed to estimate $\beta \equiv \Sigma^{-1}(\mu_1 - \mu_0)$ directly with sparsity constraints

Research Perform an analysis like that of Efron (1975) when $d \to \infty$ as well.

Ex VI: Naïve Bayes

- may feel uncomfortable with assuming $p_k(\boldsymbol{x}) \sim \mathrm{N}(\boldsymbol{\mu}_k, \boldsymbol{\Sigma})$
- data scientists have long used (and still like) the model,

$$p_k(\boldsymbol{x}) = \prod_{j=1}^d f_{k,j}(x_j)$$

where each $f_{k,j}(\cdot)$ can be estimated separately

• especially helpful if the predictors are of mixed types (e.g., some continuous, some categorical)

Ex VI: Naïve Bayes

- may feel uncomfortable with assuming independence
- but

$$\log \frac{\mathbb{P}(y=1|\boldsymbol{x})}{\mathbb{P}(y=0|\boldsymbol{x})} = \underbrace{\log \frac{\pi_1}{\pi_0}}_{\alpha} + \sum_{j=1}^d \underbrace{\log \frac{f_{1,j}(x_j)}{f_{0,j}(x_j)}}_{g_j(x_j)}$$
$$\equiv \alpha + \sum_{j=1}^d g_j(x_j),$$

and most people comfortable with generalizing linear logistic regression to additive logistic regression

Ex VII: Neural Networks

• sigmoid function

$$\sigma(u) = \frac{e^u}{1 + e^u}$$

• hidden layer $\ell=1,2,...,L\!-\!1$

$$z_t^{(\ell)} = \sigma \left(\alpha_t^{(\ell)} + \sum_b w_{t,b}^{(\ell)} z_b^{(\ell-1)} \right)$$

• top layer L

$$z_t^{(L)} = \mathbb{P}(y = t|...)$$

• bottom layer 0

$$z_b^{(0)} = x_b, \quad b = 1, 2, ..., d$$



Ex VIII: Nearest-Neighbor Classifier

• can also estimate $\mathbb{P}(y|\boldsymbol{x})$ by

 $\widehat{\mathbb{P}}(y = k | \boldsymbol{x}) = \operatorname{fraction} \{ y_i = k : \boldsymbol{x}_i \in \mathcal{N}(\boldsymbol{x}) \},\$

where $\mathcal{N}(\boldsymbol{x})$ denotes a "neighborhood" around \boldsymbol{x}

Bayes Error

Myth If my misclassification error is not very close to zero, my classifier must not be very good.

Myth If I know the true model, my misclassification error must be zero.

Truth Even if we knew $\mathbb{P}(y|\boldsymbol{x})$ (or $p_k(\boldsymbol{x}), \pi_k, ...$) perfectly, we might still have considerable misclassification error — these errors are called the Bayes error.



Summary

- key ideas:
 - regression; mean squared error; $\mathbb{E}(y|\boldsymbol{x})$
 - bias-variance trade-off; curse of dimensionality
 - classification; mean 0-1 error; $\mathbb{P}(y|\boldsymbol{x})$; Bayes error
- specific methods:
 - linear regression; nearest-neighbors; kernel regression
 - logistic regression; linear discriminant analysis
 - naïve Bayes; neural network
 - graphical LASSO
- didn't discuss:
 - actual estimation procedures

